# MLCC 2018 Dimensionality Reduction and PCA

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#### **Outline**

PCA & Reconstruction

PCA and Maximum Variance

PCA and Associated Eigenproblem

Beyond the First Principal Component

PCA and Singular Value Decomposition

Kernel PCA

## **Dimensionality Reduction**

In many practical applications it is of interest to reduce the dimensionality of the data:

- data visualization
- data exploration: for investigating the "effective" dimensionality of the data

## **Dimensionality Reduction (cont.)**

This problem of dimensionality reduction can be seen as the problem of defining a map

$$M: X = \mathbb{R}^D \to \mathbb{R}^k, \quad k \ll D,$$

according to some suitable criterion.

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In the following data reconstruction will be our guiding principle.

# **Principal Component Analysis**

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PCA can be derived from several prospective and here we give a **geometric** derivation.

# **Dimensionality Reduction by Reconstruction**

Recall that, if

$$w \in \mathbb{R}^D$$
,  $||w|| = 1$ ,

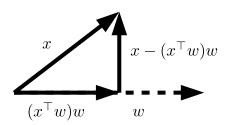
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# Dimensionality Reduction by Reconstruction (cont.)

First, consider k = 1. The associated **reconstruction error** is

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#### Problem:

Find the direction p allowing the best reconstruction of the training set

# Dimensionality Reduction by Reconstruction (cont.)



Let  $\mathbb{S}^{D-1} = \{w \in \mathbb{R}^D \mid ||w|| = 1\}$  is the sphere in D dimensions. Consider the **empirical reconstruction** minimization problem,

$$\min_{w \in \mathbb{S}^{D-1}} \frac{1}{n} \sum_{i=1}^{n} \|x_i - (w^T x_i) w\|^2.$$

The solution p to the above problem is called the **first principal** component of the data

## **An Equivalent Formulation**

A direct computation shows that  $\|x_i - (w^Tx_i)w\|^2 = \|x_i\|^2 - (w^Tx_i)^2$ 

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Then, problem

$$\min_{w \in \mathbb{S}^{D-1}} \frac{1}{n} \sum_{i=1}^{n} \|x_i - (w^T x_i) w\|^2$$

is equivalent to

$$\max_{w \in \mathbb{S}^{D-1}} \frac{1}{n} \sum_{i=1}^{n} (w^T x_i)^2$$

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#### **Reconstruction and Variance**

Assume the data to be centered,  $\bar{x} = \frac{1}{n}x_i = 0$ , then we can interpret the term

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as the **variance** of x in the direction w.

The first PC can be seen as the direction along which the data have maximum variance.

$$\max_{w \in \mathbb{S}^{D-1}} \frac{1}{n} \sum_{i=1}^{n} (w^T x_i)^2$$

# **Centering**

If the data are not centered, we should consider

$$\max_{w \in \mathbb{S}^{D-1}} \frac{1}{n} \sum_{i=1}^{n} (w^{T}(x_{i} - \bar{x}))^{2}$$
 (1)

equivalent to

$$\max_{w \in \mathbb{S}^{D-1}} \frac{1}{n} \sum_{i=1}^{n} (w^T x_i^c)^2$$

with  $x^c = x - \bar{x}$ .

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## **Centering and Reconstruction**

If we consider the effect of centering to reconstruction it is easy to see that we get

$$\min_{w,b \in \mathbb{S}^{D-1}} \frac{1}{n} \sum_{i=1}^{n} ||x_i - ((w^T(x_i - b))w + b)||^2$$

where

$$((w^T(x_i-b))w+b$$

is an affine (rather than an orthogonal) projection

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Using the symmetry of the inner product,

$$\frac{1}{n} \sum_{i=1}^{n} (w^{T} x_{i})^{2} = \frac{1}{n} \sum_{i=1}^{n} w^{T} x_{i} w^{T} x_{i} = \frac{1}{n} \sum_{i=1}^{n} w^{T} x_{i} x_{i}^{T} w = w^{T} (\frac{1}{n} \sum_{i=1}^{n} x_{i} x_{i}^{T}) w$$

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Then, we can consider the problem

$$\max_{w \in \mathbb{S}^{D-1}} w^T C_n w, \quad C_n = \frac{1}{n} \sum_{i=1}^n x_i x_i^T$$

We make two observations:

▶ The ("covariance") matrix  $C_n = \frac{1}{n} \sum_{i=1}^n X_n^T X_n$  is symmetric and positive semi-definite.

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Note that, if  $C_n u = \lambda u$  then  $\frac{u^T C_n u}{u^T u} = \lambda$ , since u is normalized.

Indeed, it is possible to show that the Rayleigh quotient achieves its maximum at a vector corresponding to the maximum eigenvalue of  $C_n$ 

Computing the first principal component of the data reduces to computing the biggest eigenvalue of the covariance and the corresponding eigenvector.

$$C_n u = \lambda u, \quad C_n = \frac{1}{n} \sum_{i=1}^n X_n^T X_n$$

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## **Beyond the First Principal Component**

We discuss how to consider more than one principle component  $\left(k>1\right)$ 

$$M: X = \mathbb{R}^D \to \mathbb{R}^k, \quad k \ll D$$

The idea is simply to iterate the previous reasoning

#### **Residual Reconstruction**

The idea is to consider the one dimensional projection that can best reconstruct the residuals

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An associated minimization problem is given by

$$\min_{w \in \mathbb{S}^{D-1}, w \perp p} \frac{1}{n} \sum_{i=1}^{n} ||r_i - (w^T r_i) w||^2.$$

(note: the constraint  $w \perp p$ )

# Residual Reconstruction (cont.)

Note that for all  $i = 1, \ldots, n$ ,

$$||r_i - (w^T r_i)w||^2 = ||r_i||^2 - (w^T r_i)^2 = ||r_i||^2 - (w^T x_i)^2$$

since  $w\perp p$ 

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Then, we can consider the following equivalent problem

$$\max_{w \in \mathbb{S}^{D-1}, w \perp p} \frac{1}{n} \sum_{i=1}^{n} (w^{T} x_{i})^{2} = w^{T} C_{n} w.$$

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#### PCA as an Eigenproblem

$$\max_{w \in \mathbb{S}^{D-1}, w \perp p} \frac{1}{n} \sum_{i=1}^{n} (w^{T} x_{i})^{2} = w^{T} C_{n} w.$$

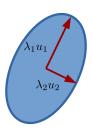
Again, we have to minimize the Rayleigh quotient of the covariance matrix with the extra constraint  $w\perp p$ 

Similarly to before, it can be proved that the solution of the above problem is given by the second eigenvector of  $\mathcal{C}_n$ , and the corresponding eigenvalue.

## PCA as an Eigenproblem (cont.)

$$C_n u = \lambda u, \quad C_n = \frac{1}{n} \sum_{i=1}^n x_i x_i^T$$

The reasoning generalizes to more than two components: computation of k principal components reduces to finding k eigenvalues and eigenvectors of  $\mathcal{C}_n$ .



#### Remarks

▶ Computational complexity roughly  $O(kD^2)$  (complexity of forming  $C_n$  is  $O(nD^2)$ ). If we have n points in D dimensions and  $n \ll D$  can we compute PCA in less than  $O(nD^2)$ ?

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#### **Remarks**

- ▶ Computational complexity roughly  $O(kD^2)$  (complexity of forming  $C_n$  is  $O(nD^2)$ ). If we have n points in D dimensions and  $n \ll D$  can we compute PCA in less than  $O(nD^2)$ ?
- ► The dimensionality reduction induced by PCA is a linear projection. Can we generalize PCA to non linear dimensionality reduction?

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# **Singular Value Decomposition**

Consider the data matrix  $X_n$ , its singular value decomposition is given by

$$X_n = U\Sigma V^T$$

#### where:

- ▶ U is a n by k orthogonal matrix,
- V is a D by k orthogonal matrix,
- $ightharpoonup \Sigma$  is a diagonal matrix such that  $\Sigma_{i,i} = \sqrt{\lambda_i}$ ,  $i = 1, \ldots, k$  and  $k \leq \min\{n, D\}$ .

The columns of U and the columns of V are the left and right singular vectors and the diagonal entries of  $\Sigma$  the singular values.

# Singular Value Decomposition (cont.)

The SVD can be equivalently described by the equations

$$C_n p_j = \lambda_j p_j, \quad \frac{1}{n} K_n u_j = \lambda_j u_j,$$
  
$$X_n p_j = \sqrt{\lambda_j} u_j, \quad \frac{1}{n} X_n^T u_j = \sqrt{\lambda_j} p_j,$$

for  $j=1,\dots,d$  and where  $C_n=\frac{1}{n}X_n^TX_n$  and  $\frac{1}{n}K_n=\frac{1}{n}X_nX_n^T$ 

# **PCA** and Singular Value Decomposition

If  $n \ll p$  we can consider the following procedure:

- form the matrix  $K_n$ , which is  $O(Dn^2)$
- find the first k eigenvectors of  $K_n$ , which is  $O(kn^2)$
- compute the principal components using

$$p_{j} = \frac{1}{\sqrt{\lambda_{j}}} X_{n}^{T} u_{j} = \frac{1}{\sqrt{\lambda_{j}}} \sum_{i=1}^{n} x_{i} u_{j}^{i}, \quad j = 1, \dots, d$$

where  $u=(u^1,\dots,u^n)$ , This is O(knD) if we consider k principal components.

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## **Beyond Linear Dimensionality Reduction?**

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...it is easy to think of situations where this assumption might violated.

Can we use kernels to obtain non linear generalization of PCA?

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#### From SVD to KPCA

Using SVD the projection of a point x on a principal component  $p_i$ , for  $j=1,\ldots,d$ , is

$$(M(x))^j = x^T p_j = \frac{1}{\sqrt{\lambda_j}} x^T X_n^T u_j = \frac{1}{\sqrt{\lambda_j}} \sum_{i=1}^n x^T x_i u_j^i,$$

Recall

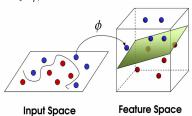
$$C_n p_j = \lambda_j p_j, \quad \frac{1}{n} K_n u_j = \lambda_j u_j,$$
$$X_n p_j = \sqrt{\lambda_j} u_j, \quad \frac{1}{n} X_n^T u_j = \sqrt{\lambda_j} p_j,$$

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### **PCA** and Feature Maps

$$(M(x))^j = \frac{1}{\sqrt{\lambda_j}} \sum_{i=1}^n x^T x_i u_i^i,$$

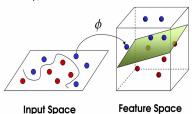
What if consider a non linear feature-map  $\Phi: X \to F$ , before performing PCA?



### **PCA** and Feature Maps

$$(M(x))^{j} = \frac{1}{\sqrt{\lambda_{j}}} \sum_{i=1}^{n} x^{T} x_{i} u_{i}^{i},$$

What if consider a non linear feature-map  $\Phi: X \to F$ , before performing PCA?



$$(M(x))^j = \Phi(x)^T p_j = \frac{1}{\sqrt{\lambda_j}} \sum_{i=1}^n \Phi(x)^T \Phi(x_i) u_j^i,$$

where  $K_n \sigma_j = \sigma_j u_j$  and  $(K_n)_{i,j} = \Phi(x)^T \Phi(x_j)$ .

#### **Kernel PCA**

$$(M(x))^j = \Phi(x)^T p_j = \frac{1}{\sqrt{\lambda_j}} \sum_{i=1}^n \Phi(x)^T \Phi(x_i) u_j^i,$$

If the feature map is defined by a positive definite kernel  $K: X \times X \to \mathbb{R}$ , then

$$(M(x))^{j} = \frac{1}{\sqrt{\lambda_{j}}} \sum_{i=1}^{n} K(x, x_{i}) u_{j}^{i},$$

where  $K_n \sigma_j = \sigma_j u_j$  and  $(K_n)_{i,j} = K(x_i, x_j)$ .

## Wrapping Up

In this class we introduced PCA as a basic tool for dimensionality reduction. We discussed computational aspect and extensions to non linear dimensionality reduction (KPCA)

#### **Next Class**

In the next class, beyond dimensionality reduction, we ask how we can devise interpretable data models, and discuss a class of methods based on the concept of sparsity.